

Modular Approaches in SAVE

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Summer School on the Design and Analysis of Computer Experiments

Simon Fraser University

August 11-16, 2006

Outline

The Problem

Full Bayes

Modular Approaches

Discussions

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The Problem

- ▶ Recall the `SAVE` model for computer model validation,

$$y^F(\cdot) = y^M(\cdot) + b(\cdot) + e(\cdot)$$

- ▶ The inputs may include,
 - ▶ \mathbf{u}^* : calibration parameters (computer model associated).
 - ▶ \mathbf{x}_1 : controllable parameters (configuration).
 - ▶ \mathbf{x}_2^* : unknown parameters varying from specimen to specimen.

$$y_i^F(\mathbf{x}_1, \mathbf{x}_2^*, t) = y^M(\mathbf{u}^*, \mathbf{x}_1, \mathbf{x}_2^*, t) + b(\mathbf{x}_1, \mathbf{x}_2^*, t) + e_i(t)$$

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The Problem

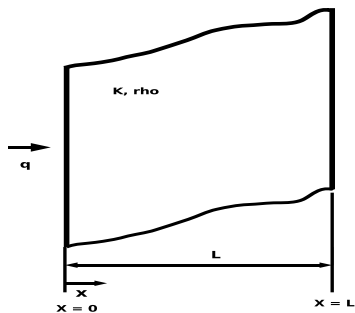
- ▶ A simplified `SAVE` model is,

$$y_i^F(\mathbf{x}_1, \mathbf{x}_2^*, t) = y^M(\mathbf{u}^*, \mathbf{x}_1, \mathbf{x}_2^*, t) + b(\mathbf{x}_1, t) + e_i^1(t)$$

where $e_i^1(\cdot)$ includes both field measurement error and a nugget to accommodate the simplification in the bias function.

Confounding may still exist between \mathbf{x}_2^* and the bias process!

The Sandia Example



- ▶ thermal conductivity $\kappa(W/m^{\circ}C)$;
- ▶ volumetric heat capacity $\rho(J/m^3{}^{\circ}C)$;
- ▶ heat flux $q(W/m^2)$;
- ▶ device thickness $L(m)$;
- ▶ initial temperature $T_0(^{\circ}C)$;
- ▶ distance to surface $x(m)$;
- ▶ time $t(s)$.

- ▶ The Thermal computer model,

$$T^M(\kappa, \rho, T_0, x, L, q; t) = T_0 + \frac{qL}{\kappa} \left(\frac{\kappa t / \rho}{L^2} + \frac{1}{3} - \frac{x}{L} + \frac{x^2}{2L^2} - \sum_1^6 \frac{2 \exp(-\frac{n^2 \pi^2 \kappa t}{L^2 \rho}) \cos(n\pi \frac{x}{L})}{\pi^2 n^2} \right)$$

- ▶ Controllable variables for the i'th specimen,

$$\mathbf{x}_1^{(i)} = (q_i, L_i, x_i, T_{i0})$$

- ▶ Unknown variables for the i'th specimen,

$$\mathbf{x}_2^{(i)} = (\kappa_i, \rho_i)$$

- ▶ We use $\mathbf{x}_2^{(i*)} = (\kappa_i^*, \rho_i^*)$ represent the associated true values.

The Bayesian Hierarchical Model

- ▶ Gaussian stochastic processes,

$$b(\cdot, \cdot) \sim \text{GP} \left(0, \tau^2 c_{x_1}(\cdot, \cdot) c_t(\cdot, \cdot) \right), \epsilon(\cdot) \sim \text{GP} \left(0, \sigma^2 c_t(\cdot, \cdot) \right)$$

- ▶ Power exponential families for the correlations,

$$c_{x_1}(\mathbf{x}_1, \mathbf{x}'_1) = \exp \left(- \sum_1^3 |\beta_k(\mathbf{x}_{1k} - \mathbf{x}'_{1k})|^{\alpha_k} \right)$$
$$c_t(t, t') = \exp(-|\beta^{(t)}(t - t')|^{\alpha^{(t)}})$$

The unknowns are $\{\kappa^*, \rho^*, \alpha, \beta, \sigma^2, \tau^2, \mathbf{b}\}$.

$$\pi(\tau^2) \propto \exp(-0.001\tau^2)$$

$$\pi(\sigma^2) \propto 1/\sigma^2$$

$$\pi(\alpha) \propto \text{U}(\alpha \mid 1, 2)$$

$$\pi(\beta) \sim \exp(\lambda\beta)$$

$$\pi(\kappa^*)$$

$$\pi(\rho^*)$$

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Full Bayes Approach

We can use MCMC algorithm to draw posterior samples for the parameters.

- ▶ Sample $(\kappa^*, \rho^* \mid \text{Data}, \alpha, \beta, \sigma^2, \tau^2)$ by Metropolis-Hastings algorithm.
- ▶ Sample $(\alpha \mid \text{Data}, \kappa^*, \rho^*, \beta, \sigma^2, \tau^2)$ by Metropolis-Hastings algorithm.
- ▶ Sample $(\beta \mid \text{Data}, \kappa^*, \rho^*, \alpha, \sigma^2, \tau^2)$ by Metropolis-Hastings algorithm.
- ▶ Sample $(\mathbf{b} \mid \text{Data}, \kappa^*, \rho^*, \alpha, \beta, \sigma^2, \tau^2)$.
- ▶ Sample $(\tau^2 \mid \text{Data}, \kappa^*, \rho^*, \alpha, \beta, \tau^2, \mathbf{b})$.
- ▶ Sample $(\sigma^2 \mid \text{Data}, \kappa^*, \rho^*, \alpha, \beta, \sigma^2, \mathbf{b})$.

Results from Full Bayes Approach

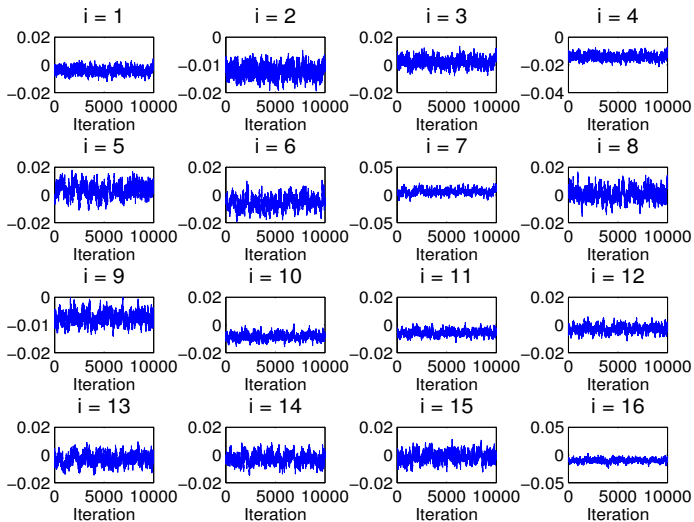


Figure: Trace plots for κ^*

Results from Full Bayes Approach

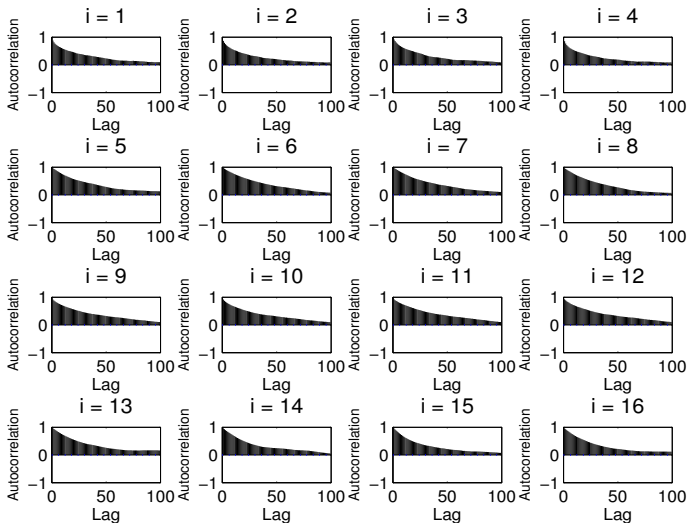


Figure: Acf plots for κ^*

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Modular Approach 1

- ▶ Unlike the full Bayes approach, modular approaches first try to fix (α, β) in the bias process.
- ▶ Modular Approach 1:
 - ▶ Get samples of α, β by first approximating the SAVE model by,

$$y_i^F(\kappa_i^*, \rho_i^*, x_i, L_i, q_i; t) \approx y^M(\hat{\kappa}, \hat{\rho}, x_i, L_i, q_i; t) + b(x_i, L_i, q_i; t) + \epsilon_i^{(2)}(t)$$

where $(\hat{\kappa}, \hat{\rho})$ are the prior means of κ^* and ρ^* .

- ▶ Fix α, β at posterior means $\hat{\alpha}, \hat{\beta}$, and draw samples for $(\kappa^*, \rho^*, \sigma^2, \tau^2, \mathbf{b} \mid \text{Data}, \hat{\alpha}, \hat{\beta})$.

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Results from Modular Approach 1

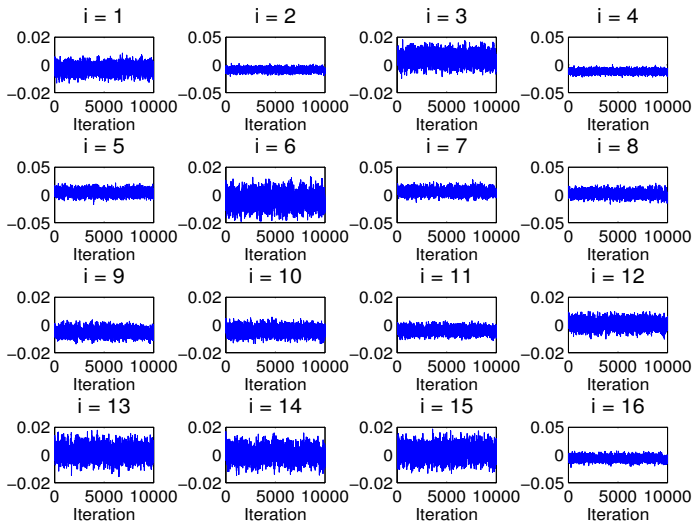


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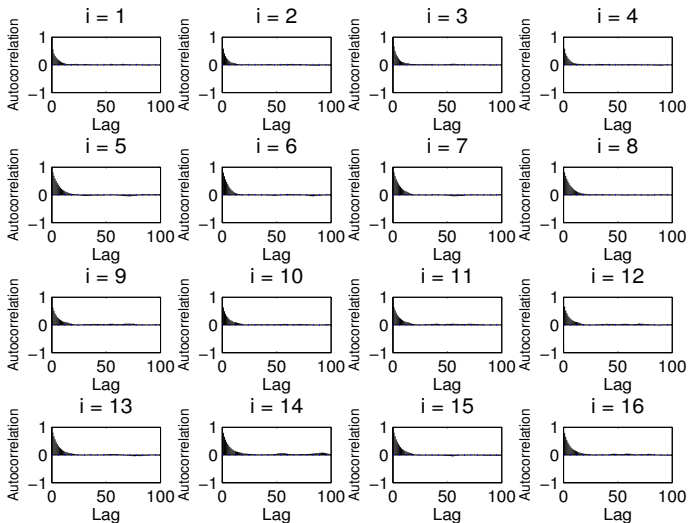


Figure: Acf plots for κ^*

Modular Approach 2

Modular Approach 2:

- ▶ Draw samples for α, β by first approximating the SAVE model by,

$$y_i^F(\kappa_i^*, \rho_i^*, \mathbf{x}_i, L_i, \mathbf{q}_i; t) \approx \mathbb{E} \left(y^M(\kappa_i^*, \rho_i^*, \mathbf{x}_i, L_i, \mathbf{q}_i; t) \right) \\ + b(\mathbf{x}_i, L_i, \mathbf{q}_i; t) + \epsilon_i^{(3)}(t)$$

This can be done by sampling $\kappa_i^* \sim \pi(\kappa_i^*)$, $\rho_i^* \sim \pi(\rho_i^*)$.

- ▶ Fix α, β at the posterior means $\hat{\alpha}, \hat{\beta}$, and draw samples for $(\boldsymbol{\kappa}^*, \boldsymbol{\rho}^*, \sigma^2, \tau^2, \mathbf{b} \mid \text{Data}, \hat{\alpha}, \hat{\beta})$.

Results from Modular Approach 2

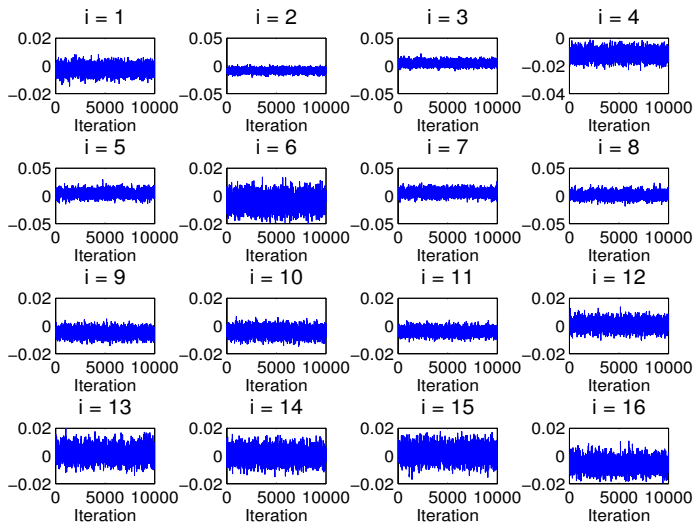


Figure: Trace plots for κ^*

Results from Modular Approach 2

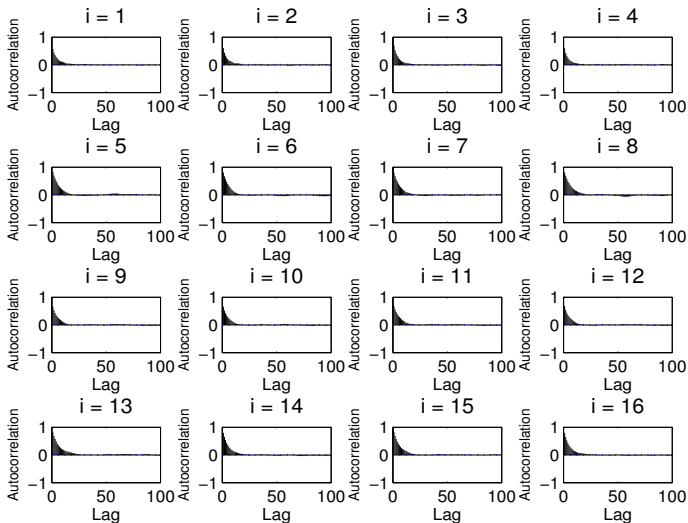


Figure: Acf plots for κ^*

Modular Approach 3

Modular Approach 3:

- ▶ Draw samples for (κ^*, ρ^*) by first assuming $\mathbf{b} = 0$ in the SAVE model, i.e.,

$$y_j^F(\kappa_j^*, \rho_j^*, x_i, L_i, \mathbf{q}_i; t) = y^M(\kappa_j^*, \rho_j^*, x_i, L_i, \mathbf{q}_i; t) + \epsilon_j^{(4)}(t)$$

- ▶ Fix κ^*, ρ^* at posterior means $\hat{\kappa}^*, \hat{\rho}^*$, and draw samples for $(\alpha, \beta, \sigma^2, \tau^2, \mathbf{b} \mid \text{Data}, \hat{\kappa}^*, \hat{\rho}^*)$.
- ▶ Fix α, β at the posterior means $\hat{\alpha}, \hat{\beta}$, and draw samples for $(\kappa^*, \rho^*, \sigma^2, \tau^2, \mathbf{b} \mid \text{Data}, \hat{\alpha}, \hat{\beta})$.

Results from Modular Approach 3

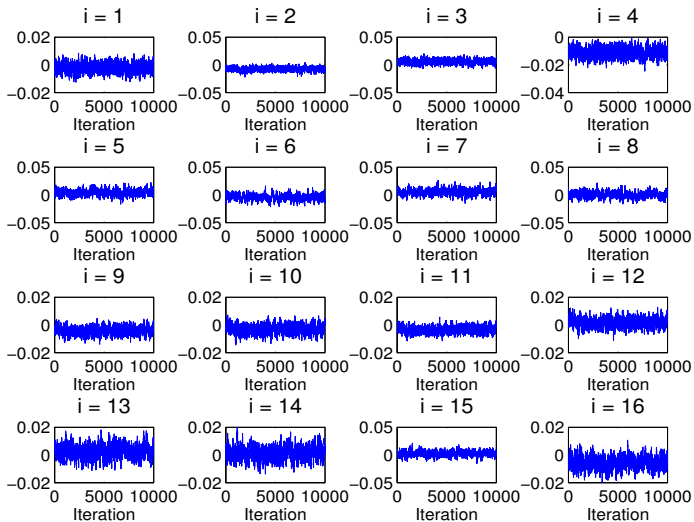


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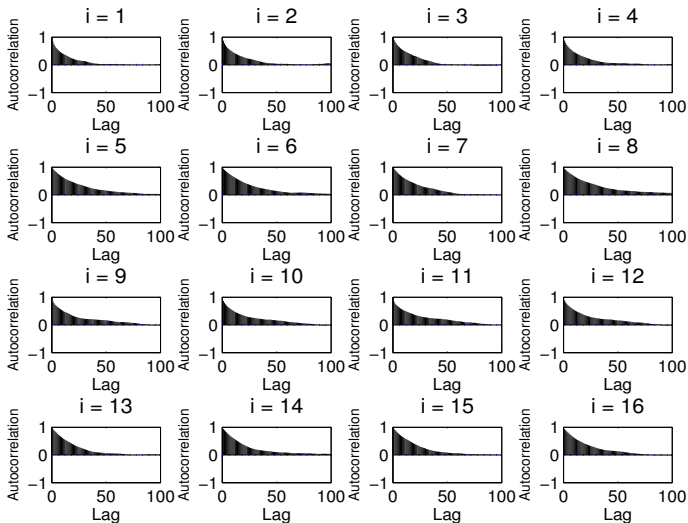


Figure: Acf plots for κ^*

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Posterior Summaries

	FB	MA 1	MA 2	MA 3
β_2	5.40 (1.56, 17.56)	10.75 (1.29, 36.86)	14.62 (1.20, 115.49)	38.71 (7.85, 120.44)
$\beta_3 \times 10^6$	1.08 (0.56, 2.82)	1.12 (0.55, 2.91)	1.07 (0.51, 2.91)	1.10 (0.56, 2.79)
$\alpha^{(t)}$	1.941 (1.898, 1.966)	1.998 (1.997, 1.999)	1.998 (1.997, 1.999)	1.9985 (1.9975, 1.9991)
$\beta^{(t)} \times 10^3$	1.20 (0.93, 1.54)	1.06 (0.94, 1.19)	1.05 (0.94, 1.19)	0.94 (0.82, 1.06)
τ^2	2079 (1209, 3755)	23307 (14940, 37647)	22843 (14658, 36737)	30055 (18889, 48386)
σ^2	1.55 (0.98, 2.57)	41.55 (32.37, 54.39)	43.25 (33.77, 57.27)	65.74 (50.95, 86.40)

Table: Posterior medians with 95% predictive intervals

The Bias functions

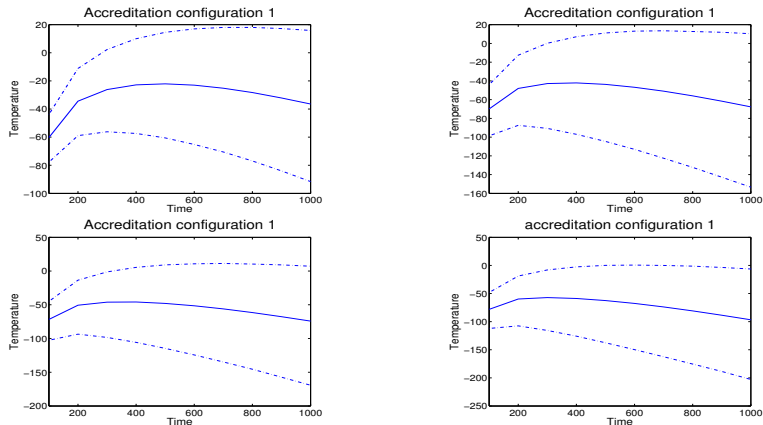


Figure: The bias functions for: upper-left: full Bayes approach; upper-right: approach 1; lower-left: approach 2; lower-right: approach 3

Pure Model predictions

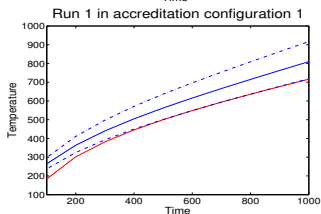
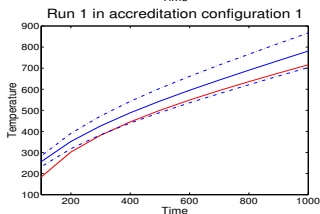
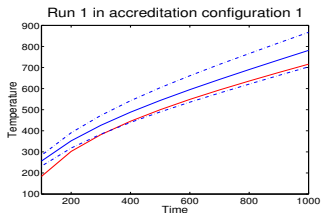
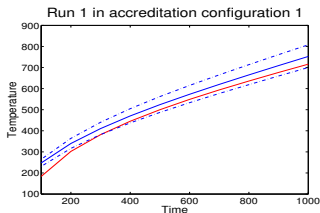


Figure: The pure model predictions for upper-left: full Bayes approach; upper-right: approach 1; lower-left: approach 2; lower-right: approach 3

Predictions for the reality

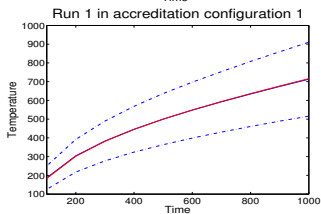
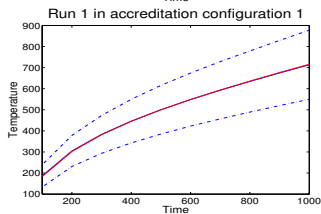
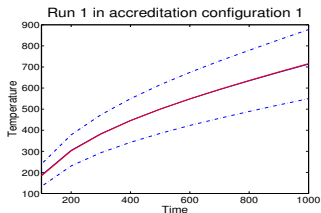
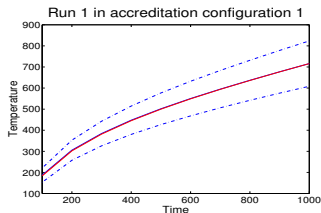


Figure: The predictions for the reality — upper-left: full Bayes approach; upper-right: approach 1; lower-left: approach 2; lower-right: approach 3

Conclusion

- ▶ The range and roughness parameters are similar in all the approaches.
- ▶ The Full Bayes approach has smaller variances both for the bias processes and the nuggets. But the auto-correlation functions of κ^* is suggesting that it does not converge.
- ▶ The modular approach 1 and 2 have bigger variances both for the bias processes and the nuggets. But the auto-correlation functions shows that the samples of κ^* are less correlated.
- ▶ Modular approach 1 allows fast computation (use of MLE).
- ▶ The modular approach 3 does not solve the confounding problem. Acf decays slowly.

Thank you!