

## Random Variables

Let  $\Omega$  be any set,  $\mathcal{F}$  any Sigma Field on  $\Omega$ , and  $\mathbb{P}$  any probability measure defined for each element of  $\mathcal{F}$ ; such a triple  $(\Omega, \mathcal{F}, \mathbb{P})$  is called a *probability space*. Let  $\mathbb{R}$  denote the real numbers  $(-\infty, \infty)$  and  $\mathcal{B}$  the Borel sets on  $\mathbb{R}$  generated by (for example) the half-open sets  $(a, b]$ .

**Definition.** A real-valued Random Variable is a function  $X : \Omega \rightarrow \mathbb{R}$  that is “ $\mathcal{F} \setminus \mathcal{B}$ -measurable,” *i.e.*, that satisfies  $X^{-1}(B) = \{\omega : X(\omega) \in B\} \in \mathcal{F}$  for each Borel set  $B \in \mathcal{B}$  (or, equivalently, simply for each set  $B$  of the form  $(-\infty, b]$  for some rational  $-\infty < b < \infty$ ).

This is sometimes denoted simply “ $X^{-1}(\mathcal{B}) \subset \mathcal{F}$ .” Since the probability measure  $\mathbb{P}$  is only defined on sets  $F \in \mathcal{F}$ , a random variable *must* satisfy this condition if we are to be able to find the probability  $\Pr[X \in B]$  for each Borel set  $B$ , or even if we want to find the distribution function (DF)  $F_X(b) \equiv \Pr[X \leq b]$  for each rational number  $b$ . Note that set-inverses are rather well-behaved functions from one class of sets to another; specifically, for any collection  $\{A_\alpha\} \subset \mathcal{B}$ ,

$$[X^{-1}(A_\alpha)]^c = X^{-1}((A_\alpha)^c) \quad \bigcap_{\alpha} X^{-1}(A_\alpha) = X^{-1}\left(\bigcap_{\alpha} A_\alpha\right) \quad \bigcup_{\alpha} X^{-1}(A_\alpha) = X^{-1}\left(\bigcup_{\alpha} A_\alpha\right)$$

and thus, measurable or not,  $X^{-1}(\mathcal{B})$  is a Sigma Field if  $\mathcal{B}$  is; it is denoted  $\mathcal{F}_X$  (or  $\sigma(X)$ ), called the “sigma field generated by  $X$ ,” and is the smallest sigma field  $\mathcal{G}$  such that  $X$  is  $(\mathcal{G} \setminus \mathcal{B})$ -measurable. In particular,  $X$  is  $(\mathcal{F} \setminus \mathcal{B})$ -measurable if and only if  $\sigma(X) \subset \mathcal{F}$ .

In probability and statistics, sigma field’s represent *information*: a random variable  $Y$  is measurable over  $\mathcal{F}_X$  if and only if the value of  $Y$  can be found from that of  $X$ , *i.e.*, if there exists some function  $\varphi$  such that  $Y = \varphi(X)$ . Note the difference in perspective between real analysis, on the one hand, and probability/statistics, on the other; in analysis it is only *Lebesgue* measurability that mathematicians worry about, and only to avoid paradoxes and pathologies. In probability and statistics we study measurability for a variety of sigma field’s, and the (technical) concept of measurability corresponds to the (empirical) notion of *observability*.

### DISTRIBUTIONS.

A random variable  $X$  on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  induces a measure  $\mu_X$  on  $(\mathbb{R}, \mathcal{B})$ , called the *distribution measure* (or simply the *distribution*), via the relation

$$\mu(B) = \mathbb{P}[X \in B],$$

sometimes written more succinctly as  $\mu_X = \mathbb{P} \circ X^{-1}$  or even  $\mathbb{P}X^{-1}$ .

### Functions of Random Variables

Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space,  $X$  a (real-valued) random variable, and  $f : \mathbb{R} \rightarrow \mathbb{R}$  a (real-valued  $\mathcal{B} \setminus \mathcal{B}$ ) measurable function. Then  $Y = f(X)$  is a random variable, *i.e.*,

$$Y^{-1}(B) = X^{-1}(f^{-1}(B)) \in \mathcal{F}$$

for any  $B \in \mathcal{B}$ . Also every continuous or piecewise-continuous real-valued function on  $\mathbb{R}$  is  $\mathcal{B} \setminus \mathcal{B}$ -measurable.

### Random Vectors

Denote by  $\mathbb{R}^2$  the set of points  $(x, y)$  in the plane, and by  $\mathcal{B}^2$  the sigma field generated by rectangles of the form  $\{(x, y) : a < x \leq b, c < y \leq d\} = (a, b] \times (c, d]$ . Note that finite unions of those rectangles form a field  $\mathcal{F}_0^2$ , so the minimal sigma field and minimal  $\lambda$  system containing  $\mathcal{F}_0^2$  coincide, and the assignment  $\lambda_0^2((a, b] \times (c, d]) = (b - a) \times (d - c)$  has a unique extension to a measure on all of  $\mathcal{B}^2$ , called two-dimensional Lebesgue measure (and denoted  $\lambda^2$ ). Of course, it's just the area of sets in the plane.

A  $\mathcal{F} \setminus \mathbb{R}^2$ -measurable mapping  $X : \Omega \rightarrow \mathbb{R}^2$  is called a (two-dimensional) *random vector*, or simply an  $\mathbb{R}^2$ -valued random variable, or (a bit ambiguously) an  $\mathbb{R}^2$ -RV. It's easy to show that the components  $X_1, X_2$  of a  $\mathbb{R}^2$ -RV  $X$  are each RV's, and conversely that for any two random variables  $X_1$  and  $X_2$  the two-dimensional RV  $(X, Y) : \Omega \rightarrow \mathbb{R}^2$  is  $\mathcal{F} \setminus \mathbb{R}^2$ -measurable, *i.e.*, is a  $\mathbb{R}^2$ -RV.

Also, any measurable (and in particular, any piecewise-continuous) function  $f : \mathbb{R}^2 \rightarrow \mathbb{R}$  induces a random variable  $f(X, Y)$ : this shows that such combinations as  $X + Y, X/Y, X \wedge Y, X \vee Y$ , *etc.* are all measurable random variables.

The same ideas work in any finite number of dimensions, so without any special notice we will regard  $n$ -tuples  $(X_1, \dots, X_n)$  as  $\mathbb{R}^n$ -valued RV's, or  $\mathcal{F} \setminus \mathcal{B}^n$ -measurable functions, and will use Lebesgue  $n$ -dimensional measure  $\lambda^n$  on  $\mathcal{B}^n$ . Again  $\sum_i X_i, \prod_i X_i, \min_i X_i$ , and  $\max_i X_i$  are all random variables.

Even if we have *infinitely many* random variables we can verify the measurability of  $\sum_i X_i, \inf_i X_i$ , and  $\sup_i X_i$ , and of  $\liminf_i X_i$ , and  $\limsup_i X_i$  as well: for example,

$$\begin{aligned} [\omega : \sup_i X_i(\omega) \leq r] &= \bigcap_{i=1}^{\infty} [\omega : X_i(\omega) \leq r] \\ [\omega : \limsup_i X_i(\omega) \leq r] &= \bigcup_{i=1}^{\infty} \bigcap_{j=i}^{\infty} [\omega : X_j(\omega) \leq r]. \end{aligned}$$

The event " $X_i$  converges" is the same as

$$[\omega : \limsup_i X_i(\omega) - \liminf_i X_i(\omega) = 0],$$

and so is  $\mathcal{F}$ -measurable and has a well defined probability  $P[\limsup_i X_i = \liminf_i X_i]$ . This is one point where countable additivity (and not just finite additivity) of  $P$  is crucial, and where  $\mathcal{F}$  needs to be a sigma field (and not just a field).

#### Example: Discrete RV's

If an RV  $X$  can take on only a finite or countable set of values, say  $b_i$ , then each set  $\Lambda_i = [\omega : X(\omega) = b_i]$  must be in  $\mathcal{F}$ , the  $\Lambda_i$  are disjoint, and  $X$  can be represented in the form

$$\begin{aligned} X(\omega) &= \sum_i b_i 1_{\Lambda_i}(\omega), \quad \text{where} \tag{*} \\ 1_{\Lambda}(\omega) &= \begin{cases} 1 & \text{if } \omega \in \Lambda \\ 0 & \text{if } \omega \notin \Lambda \end{cases} \end{aligned}$$

is the so-called *indicator function* of  $\Lambda$ . By including a term with  $b_i = 0$ , if necessary, we can assume that  $\Omega = \cup \Lambda_i$  so the  $\{\Lambda_i\}$  form a "countable partition" of  $\Omega$ . Any RV can be approximated as well as we like by a simple RV of the form (\*).

**EXPLICIT CONSTRUCTION OF SIGMA FIELDS (OMIT ON FIRST READING)****Ordinals and Transfinite Induction**

Every finite set  $S$  (say, with  $n < \infty$  elements) can be *totally ordered*  $a_1 \prec a_2 \prec a_3 \prec \dots$  in  $n!$  ways, but in some sense every one of these is the same— if  $\prec_1$  and  $\prec_2$  are two orderings, there exists a 1–1 order-preserving isomorphism  $\varphi : (S, \prec_1) \longleftrightarrow (S, \prec_2)$ . Thus *up to isomorphism* there is only one ordering for any finite set.

For countably infinite sets there are many different orderings. The obvious one is  $a_1 \prec a_2 \prec a_3 \prec \dots$ , ordered just like the positive integers  $\mathbb{N}$ ; this ordering is called  $\omega$ , the first *limit ordinal*. But we could pick any element (say,  $b_1 \in S$ ) and order the remainder of  $S$  in the usual way, but declare  $a_n \prec b_1$  for every  $n \in \mathbb{N}$ ; one element is “bigger” (in the ordering) than all the others. This is *not* isomorphic to  $\omega$ , and it is called  $\omega+1$ , the *successor* to  $\omega$ . If we set aside two elements (say,  $b_1 \prec b_2$ ) to follow all the others we have  $\omega+2$ , and similarly we have  $\omega+n$  for each  $n \in \mathbb{N}$ . The limit of all these is  $\omega+\omega$ , or  $2\omega\dots$  it is the ordering we would get if we lexicographically ordered the set  $\{(i, j) : i = 1, 2, j \in \mathbb{N}\}$  of the first two rows of integers in the first quadrant, declaring  $(1, i) \prec (2, j)$  for every  $i, j$  and otherwise  $(i, j) \prec (i, k)$  if  $j < k$ .

We would get the successor to this,  $2\omega+1$ , by extending the lexicographical ordering as we add  $(3, 1)$  to  $S$ ; in an obvious way we get  $2\omega+n$  and eventually the limit ordinals  $3\omega, 4\omega, \text{etc.}$ , and the successor ordinals  $m\omega+n$ . The limit of all these is  $\omega\omega$  or  $\omega^2$ , the lexicographical ordering of the entire first quadrant of integers  $(i, j)$ . It too has successors  $\omega^2+n$  (graphically you can think about integer triplets  $(i, j, k)$ ), and limits like  $\omega^2+\omega$  and  $\omega^3$  and  $\omega^\omega$  (which turns out to be the same as  $2^\omega$ ).

In general an ordinal is a *successor* ordinal if it has a maximal element, and otherwise is a *limit* ordinal. Every ordinal  $\alpha$  has a successor  $\alpha+1$ , and every set of ordinals  $\{\alpha_n\}$  has a limit (least upper bound)  $\lambda$ . Let  $\Omega$  be the first *uncountable* ordinal.

Proofs and constructions by transfinite induction usually have one step at each successor ordinal, and another at each limit ordinal. The *Borel sets* can be defined by transfinite construction as follows. Let  $\mathcal{F}_1$  be any class of subsets of some probability space  $\mathcal{X}$  (perhaps  $\mathcal{F}_1$  is the open sets in  $\mathcal{X} = \mathbb{R}$ , for example).

Succ: For any ordinal  $\alpha$ , let  $\mathcal{F}_{\alpha+1}$  be the class of countable unions of sets  $E_n \in \mathcal{F}_\alpha$  and their complements  $E_n^c : E_n^c \in \mathcal{F}_\alpha$ .

Lim: For any limit ordinal  $\lambda$ , let  $\mathcal{F}_\lambda = \cup_{\alpha < \lambda} \mathcal{F}_\alpha$ .

Together these define  $\mathcal{F}_\alpha$  for all ordinals, limit and successor; the sigma field *generated by*  $\mathcal{F}_1$  is just  $\mathcal{F}_\Omega$ . It remains to prove that:

1.  $\mathcal{F}_1 \subset \mathcal{F}_\Omega$ , *i.e.*,  $\mathcal{F}_\Omega$  contains the open sets;
2.  $E \in \mathcal{F}_\Omega \implies E^c \in \mathcal{F}_\Omega$ , *i.e.*,  $\mathcal{F}_\Omega$  is closed under complements;
3.  $E_n \in \mathcal{F}_\Omega \implies \cup_{n=1}^\infty E_n \in \mathcal{F}_\Omega$ , *i.e.*,  $\mathcal{F}_\Omega$  is closed under countable unions;
4.  $\mathcal{F}_\Omega \subset \mathcal{G}$  for any sigma field  $\mathcal{G}$  containing  $\mathcal{F}_1$ .

Item 1. is trivial since  $\mathcal{F}_\Omega = \cup_{\alpha < \Omega} \mathcal{F}_\alpha$ , and in particular contains  $\mathcal{F}_1$ . Item 2. follows by transfinite induction upon noting that  $E \in \mathcal{F}_\alpha \implies E^c \in \mathcal{F}_{\alpha+1}$ . Item 3 follows by noting that  $E_n \in \mathcal{F}_\Omega \implies E_n \in \mathcal{F}_{\alpha_n}$  for some  $\alpha_n < \Omega$ , and  $\beta = \sup_{n < \infty} \alpha_n$  is an ordinal satisfying  $\alpha_n \preceq \beta < \Omega$  and hence  $E_n \in \mathcal{F}_\beta$  for all  $n$  and  $\cup_{n=1}^\infty E_n \in \mathcal{F}_{\beta+1}$ . Verifying the minimality condition Item 4 is left as an exercise.

It isn't immediately obvious from the construction that we couldn't have stopped earlier—for example, that  $\mathcal{F}_2$  or  $\mathcal{F}_\omega$  isn't already the Borel sets, unchanging as we allow successively more intersections and unions. In fact that happens if the original space  $\mathcal{X}$  is countable or finite; in the case of  $\mathbb{R}$ , however, one can show that  $\mathcal{F}_\alpha \neq \mathcal{F}_{\alpha+1}$  for every  $\alpha < \Omega$ .

Do you think this explicit construction is clearer or more complicated than the completion argument used in Billingsly's book?

## INFINITE COIN TOSS

For each  $\omega \in \Omega = (0, 1]$  and integer  $n \in \mathbb{N}$  let  $\delta_n(\omega)$  be the  $n^{\text{th}}$  bit in the nonterminating binary expansion of  $\omega$ . There's some ambiguity in the dyadic expansion of rationals... for example, one-half can be written either as  $0.10b$  or as the infinitely repeating  $0.0111111...b$ . If we had used the convention that the dyadic rationals have only finitely many 1's in their expansion (so  $1/2 = 0.10b$ ) then  $\delta_n(\omega) = \lfloor 2^n \omega \rfloor \pmod{2}$ ; with our convention that all expansions must have infinitely many ones, we have

$$\delta_n(\omega) = (\lceil 2^n \omega \rceil + 1) \pmod{2}.$$

We can think of  $\{\delta_n\}$  as an infinite sequence of *random variables*, all defined on the same measurable space  $(\Omega, \mathcal{B}^1)$ , with the random variable  $\delta_1$  equal to zero on  $(0, 1/2]$  and one on  $(1/2, 1]$ ;  $\delta_2$  equal to zero on  $(0, 1/4] \cup (1/2, 3/4]$  and one on  $(1/4, 1/2] \cup (3/4, 1]$ ; and, in general,  $\delta_n$  equal to one on a union of  $2^{n-1}$  intervals, each of length  $2^{-n}$  (for a total length of  $1/2$ ), and equal to zero on the complementary set, also of length  $1/2$ . For the Lebesgue probability measure  $\mathbb{P}$  on  $\Omega$  that just assigns to each event  $E \in \mathcal{B}^1$  its length  $\mathbb{P}(E)$ , we have  $\mathbb{P}[X_n = 0] = \mathbb{P}[X_n = 1] = 1/2$ .

**Question 1:** If we had used the other convention that every binary expansion must have infinitely many zero's (instead of one's), so e.g.  $1/2 = 0.10b$ , then what would the event  $E_1 \equiv \{\omega : \delta_1(\omega) = 1\}$  have been?

The sigma field “generated by” any family of random variables  $\{X_\alpha\}$  (whether countable or not) is defined to be the smallest sigma field for which each  $X_\alpha$  is measurable, *i.e.*, the smallest one containing each  $X_\alpha^{-1}(B)$  for every Borel set  $B \subset \mathbb{R}$ . For each fixed  $n$  the  $\sigma$ -algebra  $\mathcal{F}_n$  generated by  $\delta_1, \dots, \delta_n$  is just the field  $\mathcal{F}_n = \{\cup_i (a_i/2^n, b_i/2^n)\}$  consisting of all (finite) unions of left-open intervals with both endpoints an integer over  $2^n$ . Each set in  $\mathcal{F}_n$  can be specified by listing which of the  $2^n$  intervals  $(\frac{i}{2^n}, \frac{i+1}{2^n}]$  ( $0 \leq i < 2^n$ ) it contains, so there are  $2^{2^n}$  sets in  $\mathcal{F}_n$  altogether. The union  $\cup \mathcal{F}_n$  consists of all finite unions of left-open intervals with dyadic rational endpoints. It is closed under taking complements but it still isn't a sigma field, since it isn't closed under taking *countable* unions and intersections; for example, it contains the set  $E_n = \{\omega : \delta_n = 1\}$  for each  $n \in \mathbb{N}$  and finite intersections like  $E_1 \cap \dots \cap E_n = (1 - 2^{-n}, 1]$ , but not the countable intersection  $\cap_{n=1}^\infty E_n = \{1\}$ . By definition the “join”  $\mathcal{F} = \bigvee_n \mathcal{F}_n \equiv \sigma(\cup_n \mathcal{F}_n)$  is just the smallest sigma field that contains each  $\mathcal{F}_n$  (and so contains their union); this is just the familiar Borel sets in  $(0, 1]$ .

Lebesgue measure  $\mathbb{P}$ , which assigns to any interval  $(a, b]$  its length, is determined on each  $\mathcal{F}_n$  by the rule  $\mathbb{P}[\cup_i (a_i/2^n, b_i/2^n)] = \sum (b_i - a_i)2^{-n}$  or, equivalently, by the joint distribution of the random variables  $\delta_1, \dots, \delta_n$ : independent Bernoulli's, each with  $\mathbb{P}[\delta_i = 1] = 1/2$ . For any number  $0 < p < 1$  we can make a similar measure  $\mathbb{P}_p$  on  $(\Omega, \mathcal{F}_n)$  by requiring  $\mathbb{P}_p[\delta_n = 1] = p$  and, more generally,

$$\mathbb{P}[\delta_i = d_i, 1 \leq i \leq n] = p^{\sum d_i} (1-p)^{n - \sum d_i},$$

the four intervals in  $\mathcal{F}_2$  would have probabilities  $[(1-p)^2, p(1-p), p(1-p), \text{ and } p^2]$ , for example, instead of  $[1/4, 1/4, 1/4, 1/4]$ . This determines a measure on each  $\mathcal{F}_n$ , which extends uniquely to a measure  $\mathbb{P}_p$  on  $\mathcal{F} = \bigvee_n \mathcal{F}_n$ . For  $p = 1/2$  this is Lebesgue Measure, characterized by the property that  $\mathbb{P}[(a, b]] = b - a$  for each  $0 \leq a \leq b \leq 1$ , but the other  $\mathbb{P}_p$ 's are new. This example (the family  $\delta_n$  of random variables on the spaces  $(\Omega, \mathcal{F}, \mathbb{P}_p)$ ) is an important one, and lets us build other important examples.

Under each of these probability distributions all the  $\delta_n$  are both identically distributed and independent, *i.e.*,

$$\mathbb{P}[\delta_1 \in A_1, \dots, \delta_n \in A_n] = \prod_{i=1}^n \mathbb{P}[\delta_i \in A_i].$$

Any probability assignment to intervals  $(a, b] \subset \Omega$  determines *some* joint probability distribution for all the  $\{\delta_n\}$ , but typically the  $\delta_n$  will be neither independent nor identically distributed. For any DF (*i.e.*, non-decreasing right-continuous function  $F(x)$  satisfying  $F(0) = 0$  and  $F(1) = 1$ ), the prescription  $\mathbb{P}_F((a, b]) \equiv F(b) - F(a)$  determines a probability distribution on every  $\mathcal{F}_n$  that extends uniquely to  $\mathcal{F}$ , determining the joint distribution of all the  $\{\delta_n\}$ .

**Question 2:** For  $F(x) = x^2$ , are  $\delta_1$  and  $\delta_2$  identically distributed? Independent? Find the marginal probability distribution for each  $\delta_n$  under  $\mathbb{P}_F$ .

### MEASURABILITY AND OBSERVABILITY

Fix any measure  $\mathbb{P}_p$  on  $(\Omega, \mathcal{F})$  (say, Lebesgue measure  $\mathbb{P} = \mathbb{P}_{.5}$ ), and define a new sequence of random variables  $Y_n$  on  $(\Omega, \mathcal{F}, \mathbb{P})$  by

$$Y_n(\omega) = \sum_{i=1}^n (-1)^{\delta_n(\omega)} = \sum_{i=1}^n (2\delta_n(\omega) - 1),$$

the sum of  $n$  independent terms, each  $\pm 1$  with probability  $1/2$  each. This is the “symmetric random walk” (it would be asymmetric with  $\mathbb{P}_p$  for  $p \neq .5$ ), starting at the origin and moving left or right with equal probability at each step; each  $Y_n$  is  $2S_n - n$  for the binomial  $B(n, .5)$  random variable  $S_n = \sum_{i=1}^n \delta_i$ , the partial sums of the  $\delta_n$ 's.

The sigma field generated by the first  $n$   $Y_i$ 's, that generated by the first  $n$   $S_i$ 's, and that generated by the first  $n$   $\delta_i$ 's are all the same, the finite field  $\mathcal{F}_n$  of all unions of half-open intervals with endpoints of the form  $j2^{-n}$ , and a random variable  $Z$  on  $(\Omega, \mathcal{F}, \mathbb{P})$  is  $\mathcal{F}_n$ -measurable if *and only if*  $Z$  can be written as a function  $Z = \varphi_n(\delta_1, \dots, \delta_n)$  of the first  $n$   $\delta$ 's. Thus “measurability” *means* something for us— $Z$  is **measurable** over  $\mathcal{F}_n$  if and only if you can tell its value by **observing** the first  $n$  values of  $\delta_i$  (or, equivalently, of  $Y_i$  or  $S_i$ ). We'll see that a function  $Z$  on  $\Omega$  is  $\mathcal{F}$ -measurable (*i.e.*, is a random variable) if and only if you can approximate it arbitrarily well by a function of the first  $n$   $\delta_i$ 's, as  $n \rightarrow \infty$ .

### UNIFORMS, NORMALS, AND MORE

From the infinite sequence of independent random bits  $\{\delta_n\}$  we can construct as many random variables as we like of *any* distribution, all on the same space  $(\Omega, \mathcal{F}, \mathbb{P})$ , the unit interval with Lebesgue measure (length). For example, set:

$$\begin{aligned} U_1(\omega) &= \sum_{i=1}^{\infty} 2^{-i} \delta_{2^i}(\omega) & U_3(\omega) &= \sum_{i=1}^{\infty} 2^{-i} \delta_{5^i}(\omega) \\ U_2(\omega) &= \sum_{i=1}^{\infty} 2^{-i} \delta_{3^i}(\omega) & U_4(\omega) &= \sum_{i=1}^{\infty} 2^{-i} \delta_{7^i}(\omega), \end{aligned}$$

each the sum of *different* (and therefore independent) random bits; it is easy to see that  $\{U_n\}$  will be independent, uniformly distributed random variables for  $n = 1, 2, 3, 4$ , and that we could construct as many of them as we like using successive primes  $\{2, 3, 5, 7, 11, 13, \dots\}$ .

**Question 3:** Why did I use  $\delta_{2^i}$ ,  $\delta_{3^i}$ ,  $\delta_{5^i}$ ,  $\delta_{7^i}$ ? Give another choice that would have worked.

Let  $F(x)$  be any DF (right-continuous, non-decreasing function on  $\mathbb{R}$  with limits 0 and 1  $x \rightarrow -\infty$  and  $x \rightarrow +\infty$ , respectively) and define:

$$\begin{aligned} X_1(\omega) &= \inf\{x \in \mathbb{R} : F(x) \geq U_1(\omega)\} & X_3(\omega) &= \inf\{x \in \mathbb{R} : F(x) \geq U_3(\omega)\} \\ X_2(\omega) &= \inf\{x \in \mathbb{R} : F(x) \geq U_2(\omega)\} & X_4(\omega) &= \inf\{x \in \mathbb{R} : F(x) \geq U_4(\omega)\}; \end{aligned}$$

it's not hard to see or show (we'll do it in a week or so) that the  $\{X_n\}$  are independent, each with DF  $F(x) = P[X_n \leq x]$ . For example, we could take  $X_n = \Phi^{-1}(U_n)$  to get independent random variables with the standard normal distribution or  $X_n = -\log(1 - U_n)$  for the exponential distribution.

Independent normal random variables can be constructed even more efficiently via:

$$\begin{aligned} Z_1(\omega) &= \cos(U_1)\sqrt{-2 \ln U_2} & Z_3(\omega) &= \cos(U_3)\sqrt{-2 \ln U_4} \\ Z_2(\omega) &= \sin(U_1)\sqrt{-2 \ln U_2} & Z_4(\omega) &= \sin(U_3)\sqrt{-2 \ln U_4}; \end{aligned}$$

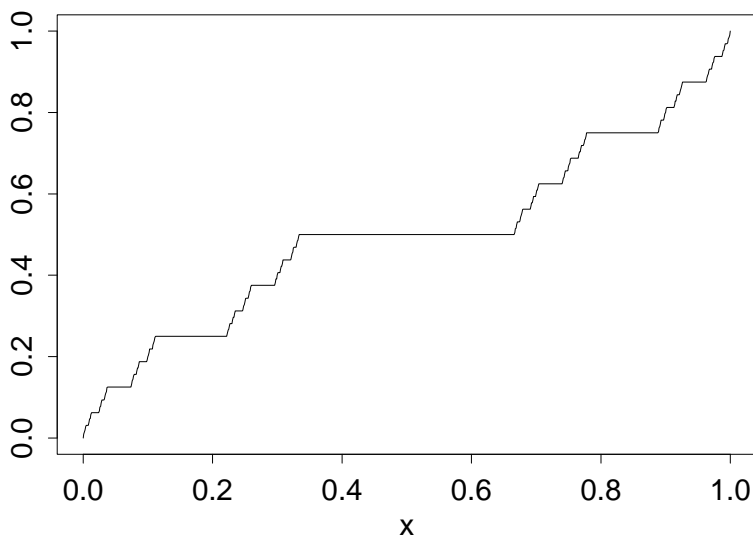
We've seen that from ordinary length measure on the unit interval (or, equivalently, from a single uniformly-distributed random variable  $\omega$ ) we can construct first an infinite sequence of independent 0 – 1 bits  $\delta_n$ ; then an infinite sequence of independent uniform random variables  $U_n$ ; then an infinite sequence of independent normal random variables  $Z_n$  or, more generally, random variables  $X_n$  with any distribution(s) we choose.

**The Cantor Distribution**

Set  $Y \equiv \sum_{n=1}^{\infty} 2\delta_n 3^{-n}$ ; then the ternery expansion of  $y = Y(\omega)$  includes only zero's (where  $\delta_n = 0$ ) and two's (where  $\delta_n = 1$ ), and so lies in the Cantor set. Since  $Y$  takes on uncountably many different values, it cannot have a discrete random variable. Its CDF can be given analytically by the expression

$$F(y) = \sum_{n=1}^{\infty} \{2^{-n} : t_n > 0, t_m \neq 1, 1 \leq m < n\},$$

in terms of the ternery expansion  $t_n \equiv [3^n y] \pmod{3}$  of  $y = \sum_{n=1}^{\infty} t_n 3^{-n}$  or graphically as



Evidently  $F(x)$  has derivative  $F' = 0$  wherever it is differentiable; this distribution is an example of a *singular* distribution, one that is neither absolutely continuous nor discrete.

**Theorem.** Let  $F(x)$  be any distribution function. Then there exist unique numbers  $p_d \geq 0$ ,  $p_c \geq 0$ ,  $p_s \geq 0$  with  $p_d + p_c + p_s = 1$  and distribution functions  $F_d(x)$ ,  $F_c(x)$ ,  $F_s(x)$  with the properties that  $F_d$  is discrete with some probability mass function  $f_d(x)$ ,  $F_c$  is absolutely

continuous with some probability density function  $f_c(x)$ , and  $F_s$  is singular, satisfying  $F(x) = p_d F_d(x) + p_c F_c(x) + p_s F_s(x)$  and

$$F_d(x) = \sum_{t \leq x} f_d(t), \quad F_c(x) = \int_{t \leq x} f_c(t) dt, \quad F'_s(x) = 0.$$

## EXPECTATION AND INTEGRAL INEQUALITIES

### Discrete RV's

If a random variable  $Y$  can take on only a finite or countably infinite set of values, say  $b_i$ , then each set  $\Lambda_i = [\omega : Y(\omega) = b_i]$  must be in  $\mathcal{F}$ ; the  $\Lambda_i$  are disjoint, and  $Y$  can be represented in the form

$$Y(\omega) = \sum_i b_i 1_{\Lambda_i}(\omega), \quad \text{where } 1_{\Lambda_i}(\omega) = \begin{cases} 1 & \text{if } \omega \in \Lambda_i \\ 0 & \text{if } \omega \notin \Lambda_i \end{cases} \quad (\star)$$

is the so-called *indicator function* of  $\Lambda_i$ . By adding a term with  $b_i = 0$ , if necessary, we can assume that  $\Omega = \cup \Lambda_i$  so the  $\{\Lambda_i\}$  form a “countable partition” of  $\Omega$ . Any RV  $X$  can be approximated as well as we like by a simple RV of the form  $(\star)$  by choosing  $\epsilon > 0$ , setting  $b_i \equiv i\epsilon$ , and

$$\Lambda_\epsilon \equiv \{\omega : b_i \leq X(\omega) < b_i + \epsilon\} \quad X_\epsilon(\omega) \equiv \sum_{-\infty}^{\infty} b_i 1_{\Lambda_i}(\omega) = \epsilon \lfloor X(\omega)/\epsilon \rfloor$$

It is easy to define the *expectation* of such a simple RV, or (equivalently) the *integral* of  $X_\epsilon$  over  $(\Omega, \mathcal{F}, \mathbf{P})$ , if  $X$  is bounded below or above (to avoid indeterminate sums):

$$\mathbf{E}X_\epsilon = \int_{\Omega} X_\epsilon(\omega) \mathbf{P}(d\omega) = \int_{\Omega} X_\epsilon(\omega) d\mathbf{P}(\omega) = \int_{\Omega} X_\epsilon d\mathbf{P} = \sum_i b_i \mathbf{P}(\Lambda_i)$$

Since  $X_\epsilon(\omega) \leq X(\omega) < X_\epsilon(\omega) + \epsilon$ , we have  $\mathbf{E}X_\epsilon \leq \mathbf{E}X < \mathbf{E}X_\epsilon + \epsilon$ , *i.e.*,

$$\sum_i i\epsilon \mathbf{P}[i\epsilon \leq X < (i+1)\epsilon] \leq \mathbf{E}X < \sum_i i\epsilon \mathbf{P}[i\epsilon \leq X < (i+1)\epsilon] + \epsilon. \quad (\star\star)$$

This determines the value of  $\mathbf{E}X = \int_{\Omega} X d\mathbf{P}$  for each random variable  $X$ . If we take  $\epsilon = 2^{-n}$  above, and simplify the notation by writing  $X_n$  for  $X_{2^{-n}} = 2^{-n} \lfloor 2^n X \rfloor$ , the sequence  $X_n$  increases monotonically to  $X$  and we can define  $\mathbf{E}X = \lim_n \mathbf{E}X_n$ .

Note that even for  $\Omega = (0, 1]$ ,  $\mathbf{P} = \lambda(dx)$  (Lebesgue measure), and  $X$  continuous, the passage to the limit suggested in  $(\star\star)$  is *not* the same as the limit of Riemann sums that is used to introduce integration in undergraduate calculus courses; for the Riemann sum it is the  $x$ -axis that is broken up into integral multiples of some  $\epsilon$ , determining the integral of *continuous* functions, while here it is the  $y$  axis that is broken up, determining the integral of all *measurable* functions. The two definitions of integral agree for continuous functions where they are both defined, of course, but the present one is much more general.

If  $X$  is *not* bounded below or above, we can set  $X^+ \equiv 0 \vee X$  and  $X^- \equiv 0 \vee -X$ , so that  $X = X^+ - X^-$  with both  $X^+$  and  $X^-$  bounded below (by zero), so their expectations are well-defined; if either  $\mathbf{E}X^+ < \infty$  or  $\mathbf{E}X^- < \infty$ , we can unambiguously define  $\mathbf{E}X \equiv \mathbf{E}X^+ - \mathbf{E}X^-$ , while if  $\mathbf{E}X^+ = \mathbf{E}X^- = \infty$  we regard  $\mathbf{E}X$  as undefined.

For any measurable set  $\Lambda \in \mathcal{F}$  we write  $\int_{\Lambda} X d\mathbf{P}$  for  $\mathbf{E}X 1_{\Lambda}$ . For  $\Omega \subset \mathbb{R}$ , if  $\mathbf{P}$  gives positive probability to either  $\{a\}$  or  $\{b\}$  then the integrals over the sets  $(a, b)$ ,  $(a, b]$ ,  $[a, b)$ , and  $[a, b]$  may all be different; the notation  $\int_a^b X d\mathbf{P}$  isn't expressive enough to distinguish them.

Frequently in Probability and Statistics we need to calculate or estimate integrals and expectations; usually this is done through limiting arguments in which a sequence of integrals is shown to converge to the one whose value we need. Here are some important properties of integrals for any measurable set  $\Lambda \in \mathcal{F}$  and random variables  $\{X_n\}$ ,  $X$ ,  $Y$ , useful for bounding or estimating the integral of a random variable  $X$  (they're only listed here for reference and so we can talk about them— don't worry, you won't have to remember them all or know how to prove them!):

1.  $\int_{\Lambda} X dP$  is well-defined and finite if and only if  $\int_{\Lambda} |X| dP < \infty$ , and  $\left| \int_{\Lambda} X dP \right| \leq \int_{\Lambda} |X| dP$ .  
We can also define  $\int_{\Lambda} X dP \leq \infty$  for any  $X$  bounded below by some  $b > -\infty$ .
2. **Lebesgue's Monotone Convergence Thm:** If  $0 \leq X_n \nearrow X$ , then  $\int_{\Lambda} X_n dP \nearrow \int_{\Lambda} X dP \leq \infty$ . In particular, the sequence of integrals converges (possibly to  $+\infty$ ).
3. **Lebesgue's Dominated Convergence Thm:** If  $X_n \rightarrow X$ , and if  $|X_n| \leq Y$  for some RV  $Y \geq 0$  with  $EY < \infty$ , then  $\int_{\Lambda} X_n dP \rightarrow \int_{\Lambda} X dP$  and  $\int_{\Lambda} |X| dP \leq \int_{\Lambda} Y dP < \infty$ . In particular, the sequence of integrals converges to a finite limit.
4. **Fatou's Lemma:** If  $X_n \geq 0$  on  $\Lambda$ , then  $\int_{\Lambda} (\liminf X_n) dP \leq \liminf (\int_{\Lambda} X_n dP)$ . The two sides may be unequal (example?), and the result is false for  $\limsup$ .
5. **Fubini's Thm:** If *either* each  $X_n \geq 0$ , *or*  $\sum_n \int_{\Lambda} |X_n| dP < \infty$ , then the order of integration and summation can be exchanged:  $\sum_n \int_{\Lambda} X_n dP = \int_{\Lambda} \sum_n X_n dP$ . If both these conditions fail, the orders may not be exchangeable (example?).
6. For any  $p > 0$ ,  $E|X|^p = \int_0^{\infty} p x^{p-1} P[|X| > x] dx$  and  $E|X|^p < \infty \Leftrightarrow \sum_{n=1}^{\infty} n^{p-1} P[|X| \geq n] < \infty$ . The case  $p = 1$  is easiest and most important: if  $S \equiv \sum_{n=1}^{\infty} P[|X| \geq n] < \infty$ , then  $S \leq E|X| < S+1$ . If  $X$  takes on only nonnegative integer values,  $EX = S$ .
7. If  $\mu_X$  is the distribution of  $X$ , and if  $f$  is a measurable real-valued function on  $\mathbb{R}$ , then  $Ef(X) = \int_{\Omega} f(X(\omega)) dP = \int_{\mathbb{R}} f(x) \mu_X(dx)$  if either side exists. In particular,  $\mu = EX = \int x \mu_X(dx)$  and  $\sigma^2 = E(X - \mu)^2 = \int (x - \mu)^2 \mu_X(dx)$ .
8. **Hölder's Inequality:** Let  $p > 1$  and  $q = \frac{p}{p-1}$  (e.g.,  $p = q = 2$  or  $p = 1.01$ ,  $q = 101$ ). Then  $EXY \leq E|XY| \leq [E|X|^p]^{\frac{1}{p}} [E|Y|^q]^{\frac{1}{q}}$ . In particular, for  $p = q = 2$ ,  
**Cauchy-Schwartz Inequality:**  $EXY \leq E|XY| \leq \sqrt{EX^2 EY^2}$ .
9. **Jensen's Inequality:** Let  $\varphi(x)$  be a convex function on  $\mathbb{R}$ ,  $X$  an integrable RV. Then  $\varphi(EX) \leq E[\varphi(X)]$ . Examples:  $\varphi(x) = |x|^p$ ,  $p \geq 1$ ;  $\varphi(x) = e^x$ ;  $\varphi(x) = [0 \vee x]$ .
10. **Markov's & Chebychev's Inequalities:** If  $\varphi$  is positive and increasing, then  $P[|X| \geq u] \leq E[\varphi(|X|)]/\varphi(u)$ . In particular  $P[|X - \mu| > u] \leq \frac{\sigma^2}{u^2}$  and  $P[|X| > u] \leq \frac{\sigma^2 + \mu^2}{u^2}$ .  
**One-Sided Version:**  $P[X > u] \leq \frac{\sigma^2}{\sigma^2 + (u - \mu)^2}$ .